Accounting for Variation in Mortality and Allee Effects When Computing Steepness for Strategic Fisheries Management ³ September 21, 2010

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13 Abstract

¹⁴ We update recent work on the scientific inference and reproductive biology of steepness in three ¹⁵ directions. First, we show how variation in natural mortality can be included in the formula ¹⁶ for steepness, for both a biomass dynamics and age-structure models. We do this using the ¹⁷ delta-method, so that only the mean and covariance of natural mortality rates appear in the ¹⁸ characterization of steepness. Second, we show how to generalize the previous methods for cases ¹⁹ in which the stock recruitment relationship is depensatory or has an Allee effect: as spawning ²⁰ population falls below a certain level, per capita reproduction starts to fall, rather than approach

a constant. We generally assume that the mechanism of depensation is imperfect fertilization 21 (and thus develop a two-sex generalization of our previous work) and determine steepness in 22 this case for both a biomass production model and an age-structured model and explore the 23 implications of such depensatory reproduction on the response of stocks to harvesting. We 24 briefly discuss how an increase in mortality as population size declines (as has been suggested 25 for penguins) could also be a mechanism for depensation. Third, we describe an improved 26 method for computing the maximum per capita reproduction in the age-structured model, and 27 show how the equivalent for the biomass dynamics model is computed. 28

²⁹ Introduction

Mangel et al (2010) developed methods for the computation of steepness of a Beverton-Holt 30 (BH) or Ricker (R) stock recruitment relationship (SRR) from first biological principles. In this 31 paper, we extend their work in two directions. First, Mangel et al (2010) show that steepness 32 can be expressed in terms of natural mortality rates and maximum per capita productivity (see 33 below for a summary). They also show that maximum per capita productivity can be estimated 34 using a stochastic simulation in which the rate of natural mortality fluctuates, drawn from a 35 known probability distribution. To improve the self-consistency of the methods, we show how 36 the probabilistic properties of natural mortality can be included, succintly approximately, in the 37 expression for steepness. 38

³⁹ Mangel *et al* (2010) also assumed that as the size of the spawning population declined the ⁴⁰ maximum per capita reproduction approached a positive constant, for example in the way that ⁴¹ $\frac{1-exp(-\alpha S)}{S}$ approaches α as $S \rightarrow 0$. A population exhibits an Allee effect or a depensatory SRR ⁴² if per capita reproduction declines as the size of the population falls below a certain level (Fowler ⁴³ and Baker 1991; Courchamp *et al* 2008, Gregory *et al* 2010). In fishery science, these effects are ⁴⁴ commonly called depensation (see Iles 1994, Liermann and Hilborn 2001, Gascoigne and Lipcius ⁴⁵ 2004 for review). Since many fish, seabird, and marine mammal populations may in principle exhibit a depensatory SRR, we show how the methods of Mangel *et al* (2010) can be extended for those kinds of SRRs. In doing so, we learn a variety of interesting qualitative properties of such systems when stressed by fishing mortality (or incidental mortality for seabirds or marine mammals). To do this, we develop an explicit two sex model for the case of a structured population. Finally, we describe an improved version of the stochastic simulation used by Mangel *et al* (2010) to compute maximum per capita reproduction.

52 Review of The Approach

⁵³ In order to make this paper self-contained, we review the approach used in Mangel *et al.* (2010).

We first consider a production or biomass dynamic model and then generalize the age-structured
 model to two sexes.

56 The Production Model

In this case, we let B(t) denote the total biomass at time t and assume that the fraction of males at birth is r, so that spawning (i.e. female) biomass at any time is (1 - r)B(t). If M is the rate of natural mortality, then in the absence of fishing the dynamics of biomass are

$$\frac{dB}{dt} = \frac{\alpha_p (1-r)B}{1+\beta B} - M \cdot B \tag{1}$$

where α_p is subscripted to note *production* and has units of new biomass per existing spawning (female) biomass per time, so that it is a rate, comparable to the rate of natural mortality M. The steady state biomass is

$$B_0 = \frac{1}{\beta} \left(\frac{\alpha_p (1-r)}{M} - 1 \right) \tag{2}$$

We thus see that existence of a steady state requires that the Beverton number (sensu Mangel 2005) $\frac{\alpha_p(1-r)}{M} > 1$ and that density dependence scales the overall size of the steady state. Thus, the parametrization in Equation 1 separates the roles of α and β in shape and scale when determining B_0 . The separation becomes even clearer when we consider steepness. ⁶⁷ According to its definition, steepness is

$$h = \frac{\frac{0.2\alpha_p(1-r)B_0}{1+0.2\beta B_0}}{\frac{\alpha_p(1-r)B_0}{1+\beta B_0}}$$
(3)

68 from which we obtain

$$h = 0.2 \cdot \frac{1 + \beta B_0}{1 + 0.2\beta B_0} \tag{4}$$

⁶⁹ However, in light of Equation 2

$$\beta B_0 = \left[\frac{\alpha_p(1-r)}{M} - 1\right] \tag{5}$$

70 so that we find

$$h = \frac{\alpha_p(1-r)}{4M + \alpha_p(1-r)} \tag{6}$$

71 which can also be rewritten as

$$h = \frac{\frac{\alpha_p(1-r)}{M}}{4 + \frac{\alpha_p(1-r)}{M}} \tag{7}$$

If the Beverton-Holt production term in Eqn 1 is replaced by the Ricker form, $\alpha_p(1-r)Be^{-\beta B}$ then Eqn 7 is replaced by (see the Appendix of Mangel *et al* (2010))

$$h = 0.2 \left(\frac{\alpha}{M}\right)^{0.8} \tag{8}$$

74 The Age-Structured Model

To generalize the age-structured model of Mangel *et al* (2010) to two sexes, we let $N_m(a, t)$ and N_f(a, t) denote the number of males and females of age a at time t respectively. The spawning stock biomass at time t is

$$B_s(t) = \sum_{a=1}^{a_{max}} N_f(a, t) W_f(a) p_{f,m}(a)$$
(9)

where $W_f(a)$ is the mass of a female at age a and $p_{f,m}$ is the probability that a female of age a is mature and a_{max} is the maximum age that an individual can attain (with suitable modification of the dynamics of the last age class, we can incorporate a 'plus' group into this formulation). If we assume that egg production is proportional to biomass then the recruited class numbers are

$$N_f(0,t) = \frac{\alpha_s(1-r)B_s(t)}{1+\beta B_s(t)}$$
(10)

$$N_m(0,t) = \frac{\alpha_s r B_s(t)}{1 + \beta B_s(t)} \tag{11}$$

where α_s has units of *new individuals/spawning biomass* and is subscripted with *s* to denote that this is a structured model. The sum of Eqns 10 and 11 gives the total recruitment. As will be seen, it does not matter whether we use spawning biomass or total biomass to characterize the density dependence.

For ages a > 0, in the absence of fishing mortality we have

$$N_f(a,t) = N_f(a-1,t-1)e^{-M_f(a-1)}$$
(12)

$$N_m(a,t) = N_m(a-1,t-1)e^{-M_m(a-1)}$$
(13)

where $M_f(a-1)$ and $M_m(a-1)$ are the natural mortality rates of females and males at age a - 1. Typically, mortality rates of males are higher than those of females. Since neither of the mortality rates depend upon time, the population will reach a steady state and a stable age distribution.

We denote the steady state female biomass by \overline{B}_s . This biomass produces a steady state female recruitment

$$\overline{N}_f(0) = \frac{\alpha_s (1-r)\overline{B}_s}{1+\beta \overline{B}_s}$$
(14)

94 The steady state female biomass is

$$\overline{B_s} = \sum_{a=1}^{a_{max}} \overline{N}_f(a) W_f(a) p_{f,m}(a) = \sum_{a=1}^{a_{max}} \overline{N}_f(0) S_f(a) W_f(a) p_{f,m}(a)$$
(15)

where $S_f(a)$ is survival from age 0 to age a, i.e. $S_f(a) = \prod_{i=0}^{a-1} e^{-M_f(i)}$. We factor $\overline{N}_f(0)$ from the last expression in Eqn 15, define $\overline{W}_f = \sum_{a=1}^{a_{max}} S_f(a) W_f(a) p_{f,m}(a)$ and thus write

$$\overline{B}_s = \overline{N}_f(0)\overline{W}_f \tag{16}$$

⁹⁷ We now rewrite Eqn 14 as

$$\overline{N}_f(0) = \frac{\alpha_s (1-r) \overline{N}_f(0) \overline{W}_f}{1+\beta \overline{B}_s}$$
(17)

98 from which we conclude

$$1 + \beta \overline{B}_s = (1 - r)\alpha_s \overline{W}_f \tag{18}$$

⁹⁹ Since reproduction is assessed counting both males and females, steepness is given by

$$h = \frac{0.2\alpha_s \overline{B}_s}{1 + 0.2\beta \overline{B}_s} \cdot \frac{1 + \beta \overline{B}_s}{\alpha_s \overline{B}_s} = \frac{0.2(1 + \beta \overline{B}_s)}{1 + 0.2\beta \overline{B}_s}$$
(19)

¹⁰⁰ We now use Eqn 18 in the last expression in Eqn 19 and simplify to obtain

$$h = \frac{(1-r)\alpha_s \overline{W}_f}{4 + (1-r)\alpha_s \overline{W}_f}$$
(20)

¹⁰¹ For the case of Ricker density dependence, Eqns 12 and 13 are replaced by

$$N_f(0,t) = \alpha_s (1-r) B_s(t) e^{-\beta B_s(t)}$$
(21)

$$N_m(0,t) = \alpha_s r B_s(t) e^{-\beta B_s(t)}$$
(22)

¹⁰² Following the steps above leads to

$$h = 0.2 \left((1-r)\alpha_s \overline{W}_f \right)^{0.8} \tag{23}$$

Comparing Eqns 7 and 8 with Eqns 20 and 23 we see that the functional form is the same and that with the exception of α_s in the latter replacing α_p in the former we can make them "identical" by setting $\overline{W}_f = \frac{1}{M}$. Mangel *et al.* (2010) explicitly show the conditions under which the result for the age-structured model becomes the result for the production model.

¹⁰⁷ Imperfect Fertilization as a Mechanism for Depensation

¹⁰⁸ Both the Ricker SRR

$$R = \alpha S e^{-\beta S}$$

109 and the Beverton-Holt SRR

$$R = \frac{\alpha S}{1 + \beta S}$$

have the property that as $S \to 0$, $R \to \alpha S$. A depensatory SRR will arise if for small Srecruitment becomes sublinear. For example, Myers *et al* (1995) modify the Beverton-Holt SRR to

$$R = \frac{\alpha S^{\delta}}{1 + \beta S^{\delta}}$$

and treat δ as a parameter to be estimated, understanding that $\delta > 1$ corresponds to depensation. Morales-Bojorquez and Nevarez-Martinez (2005) modify the Shepherd (1982) SRR (itself a modification of the Beverton-Holt) to explicitly consider a threshold level below which recruitment fails

$$R = max \left[0, \frac{\alpha(S - S_c)}{1 + \beta(S - S_c)^{\delta}} \right]$$

where S_c is the critical level at which recruitment drops to 0 (also see Chen *et al* 2002). Similarly, modifying the Ricker SRR to

$$R = \alpha S^{\delta} e^{-\beta S}$$

produces the Saila-Lorda SRR (Iles 1994) and gives a dependent or relationship whenever $\delta > 1$. 119 These are *ad hoc* modifications of the SRR, which is fine if one wants to study the problem 120 are purely a statistical one. However, if one approaches this as a problem in reproductive 121 biology of fish, then the actual mechanism for depensation becomes important. For most of the 122 analysis, we assume that the mechanism of dependent of the provided of the pro 123 size declines, but also discuss increased mortality as a mechanism for depensation. Imperfect 124 fertilization could be caused by at least two factors. First, at low population sizes individuals 125 simply may not be able to find each other. Second, even in species that form very tight mating 126 aggregations, an individual male cannot fertilize an unlimited number of females due to sperm 127 limitation. Thus, in some sense all populations must experience Allee effects and the question 128 is at what level of population size the effects become important. 129

In order to make further progress, we must select a mechanism for the Allee effect. In this paper, we consider imperfect fertilization— that some eggs remain unfertilized – as the 132 mechanism. We let

$$p_f(B) = \Pr[\text{an egg is fertilized given that the biomass of the population is } B]$$
 (24)

and consider two forms the probability of fertilization. When the biomass of the population is B the biomass of males is rB; for the algebraic representation of probability of fertilization we write

$$p_f(B) = \frac{rB}{rB_c + rB} = \frac{B}{B_c + B}$$
(25)

¹³⁶ where rB_c is the male biomass at which the probability of fertilization drops to 50%.

The second functional form is motivated by the re-analysis by Rowe *et al.* (2004) of the data of Bekkevold *et al.* (2002). Rowe *et al.* (2004) conclude that the exponential asymptotic function that best describes the fertilization rate is $0.97 \cdot (1 - e^{-2.02 \cdot N_m})$ where N_m is the number of males. We let p_0 denote the probability that a focal egg is not fertilized when a single male is present. If W_m denotes the mass of a male, then the number of males in the population when biomass is *B* is $\frac{rB}{W_m}$ so that

$$p_f(B) = 1 - p_0^{\frac{rB}{W_m}} = 1 - exp \left[\frac{rB}{W_m} \cdot \log(p_0) \right] = 1 - e^{-\gamma \cdot B}$$
(26)

where $\gamma = r |log(p_0)| / W_m$. Note that $B_c \to 0$ or $\gamma \to \infty$ correspond to Allee effects becoming weaker and weaker.

¹⁴⁵ To account for Allee effect in the production model, we replace Eqn 1by

$$\frac{dB}{dt} = \frac{\alpha_p (1-r) \cdot p_f(B) \cdot B}{1+\beta B} - M \cdot B \tag{27}$$

¹⁴⁶ with a similar modification for a Ricker SRR and for the age-structured model (Eqns 10,11).

In summary, there are four cases: 1) Beverton-Holt density dependence of recruitment and algebraic probability of fertilization; 2) Ricker density-dependence of recruitment and algebraic probability of fertilization; 3) Beverton-Holt density dependence of recruitment and exponential probability of fertilization; and 4) Ricker density-dependence of recruitment and exponential probability of fertilization. Each of these cases may apply for the biomass dynamics model or the age-structured model.

Including Variation in Natural Mortality in the Equations for Steepness

¹⁵⁵ We now show how variation in natural mortality can be incorporated into the formula for ¹⁵⁶ steepness. We begin with Eqn 7 written as

$$h = \frac{\alpha}{4M + \alpha} \equiv f(M) \tag{28}$$

¹⁵⁷ When natural mortality M fluctuates, as is assumed in the stochastic simulation used to estimate ¹⁵⁸ maximum per capita reproduction, f(M) is itself a random variable. To incorporate the variation ¹⁵⁹ in M into steepness, we use the delta method (Mangel 2006) to compute the expected value of ¹⁶⁰ f(M). That is if \overline{M} and Var(M) denote the expected value and variance of M and $E_M[\cdot]$ the ¹⁶¹ expectation over M we have

$$E_M[f(M)] = E_M[f(\overline{M}) + f'(\overline{M})(M - \overline{M}) + \frac{1}{2}f''(\overline{M})(M - \overline{M})^2$$
(29)

¹⁶² In light of Eqn 28

$$f'(M) = -\frac{4\alpha}{(4M+\alpha)^2} \tag{30}$$

$$f''(M) = \frac{32\alpha}{(4M+\alpha)^3}$$
(31)

163 We thus conclude

$$E_M[h] = \frac{\alpha}{4\overline{M} + \alpha} + \frac{16\alpha}{(4\overline{M} + \alpha)^3} Var(M)$$
(32)

which allows us to incorporate the stochastic variation of M into the formula for steepness.

¹⁶⁵ We now turn to the age-structured case, for which we rewrite Eqn 20 as

$$h = \frac{\alpha \overline{W}_f}{4 + \alpha \overline{W}_f} \tag{33}$$

where $\overline{W}_f = \sum_{a=1}^{a_{max}} S_f(a) W_f(a) p_{f,m}(a)$, more explicitly written as

$$\overline{W}_f = \sum_{a=1}^{a_{max}} exp(\sum_{a'=1}^{a-1} - M(a'))W_f(a)p_{f,m}(a)$$
(34)

We recognize that each of the M(a') in this equation may have its own probability distribution. (Mathematically speaking, \overline{W}_f is a functional, since it takes a vector of mortality rates and returns a scalar. There is a large literature on functional derivatives, but we can do all that is needed here using elementary calculus). For simplicity, we use the notation $M_1 = M(1), M_2 =$ M(2) etc to denote the rate of mortality at age, $\overline{M}_a, Var(M_a)$ and $Cov(M_a, M_{a'})$ to denote the mean, variance, and covariance of the mortality rates and $\langle \overline{W}_f \rangle$ to denote the value of \overline{W}_f obtained when the means of the rates of mortality are used. The analogue of Eqn 31 is now

$$E_M[h] = \langle \overline{W}_f \rangle + \frac{1}{2} \sum_a \frac{\partial^2 h}{\partial^2 M_a} Var(M_a) + \sum_a \sum_{a' \neq a} \frac{\partial^2 h}{\partial M_a \partial M_{a'}} Cov(M_a, M_{a'})$$
(35)

¹⁷⁴ The partial derivatives of steepness with respect to the mortality rates are

$$\frac{\partial h}{\partial M_a} = \frac{\partial h}{\partial \overline{W}_f} \cdot \frac{\partial \overline{W}_f}{\partial M_a} \tag{36}$$

$$\frac{\partial^2 h}{\partial M_a \partial M_{a'}} = \left[\frac{\partial^2 h}{\partial \overline{W}_f^2} \cdot \frac{\partial \overline{W}_f}{\partial M_a} \cdot \frac{\partial \overline{W}_f}{\partial M_a'} + \frac{\partial h}{\partial \overline{W}_f} \cdot \frac{\partial^2 \overline{W}_f}{\partial M_a \partial M_a'} \right]$$
(37)

¹⁷⁵ The easiest part in the next step is to show that

$$\frac{\partial h}{\partial \overline{W}_f} = \frac{4\alpha}{(4+\alpha \overline{W}_f)^2} \tag{38}$$

$$\frac{\partial^2 h}{\partial \overline{W}_f^2} = \frac{-8\alpha^2}{(4+\alpha \overline{W}_f)^3} \tag{39}$$

¹⁷⁶ We now evaluate the derivatives of \overline{W}_f by first rewriting Eqn 34 as

$$\overline{W}_f = e^{-M_1} W_f(1) p_{f,m}(1) + e^{-M_1 - M_2} W_f(2) p_{f,m}(2) + e^{-M_1 - M_2 - M_3} W_f(3) p_{f,m}(3) + \cdots$$
(40)

¹⁷⁷ from which we can see the pattern of first derivatives

$$\frac{\partial \overline{W}_f}{\partial M_1} = -\overline{W}_f \tag{41}$$

$$\frac{\partial \overline{W}_f}{\partial M_2} = -\overline{W}_f + e^{-M_1} W_1 p_{f,m}(1) \tag{42}$$

$$\frac{\partial W_f}{\partial M_3} = -\overline{W}_f + e^{-M_1}W_1 + e^{-M_1 - M_2}W_2 p_{f,m}(2)$$
(43)

: (44)

¹⁷⁸ Similarly, the pattern of second derivatives is

$$\frac{\partial^2 \overline{W}_f}{\partial M_1^2} = -\frac{\partial \overline{W}_f}{\partial M_1} = \overline{W}_f \tag{45}$$

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$$\frac{\partial^2 W_f}{\partial M_1 \partial M_2} = \overline{W}_f - e^{-M_1} W_1 p_{f,m}(1) \tag{46}$$

$$\frac{\partial^2 \overline{W}_f}{\partial M_2^2} = \overline{W}_f \tag{47}$$

$$\frac{\partial^2 W_f}{\partial M_1 \partial M_3} = \overline{W}_f - e^{-M_1} W_1 p_{f,m}(1) - e^{-M_1 - M_2} W_2 p_{f,m}(2)$$
(48)

(49)

It is then possible to implement Eqns 32 or 35ff in the code determining steepness. We now turn
to depensation.

¹⁸¹ Steepness for a Production Model with Depensatory SRR

- ¹⁸² For each of the cases, we can find steepness by repeating the process outlined in Eqns 1-9. Since
- it is straightforward but somewhat tedious algebra, we simply give the results.
- ¹⁸⁴ Beverton-Holt density dependence and algebraic probability of fertilization
- ¹⁸⁵ In this case, the steady state biomass satisfies the algebraic equation

$$\frac{\alpha_p \cdot (1-r) \cdot \overline{B}}{(\overline{B} + B_c)(1 + \beta \cdot \overline{B})} = M \tag{50}$$

¹⁸⁶ which can easily be solved using the quadratic formula. Steepness is given by

$$h = \frac{0.2 \frac{\alpha_p(1-r)}{M} \cdot \overline{B}}{(B_c + 0.2\overline{B})(4 + \frac{\alpha_p(1-r)\overline{B}}{M(B_c+B)})}$$
(51)

As $B_c \to 0$, this equation becomes Eqn 7, as it must. Note the important differences, however, between Eqns 7 and 51: the former only inolves α and M whereas the latter includes B_c and \overline{B} .

- ¹⁹⁰ Ricker density-dependence and algebraic probability of fertilization
- ¹⁹¹ In this case the steady state biomass \overline{B} satisfies

$$\frac{\alpha_p(1-r)}{M} \cdot \frac{\overline{B}}{B_c + \overline{B}} = e^{\beta \cdot \overline{B}}$$
(52)

Although this equation does not have an analytical solution, \overline{B} can easily be found using Newton's method (Mangel 2006). Steepness is given by

$$h = \frac{0.04(B_c + \overline{B})}{B_c + 0.2\overline{B}} \left(\frac{\alpha_p(1-r)}{M} \cdot \frac{\overline{B}}{B_c + \overline{B}}\right)^{0.8}$$
(53)

As $B_c \rightarrow 0$, this equation becomes Eqn 8, as it must. Once again, note the important differences

between Eqns 8 and 53, since the latter involves both B_c and \overline{B} .

¹⁹⁶ Beverton-Holt density dependence and exponential probability of fertilization

¹⁹⁷ In this case the steady state biomass satisfies

$$\beta \cdot \overline{B} = \frac{\alpha_p (1-r)}{M} (1 - e^{-\gamma \cdot \overline{B}}) - 1$$
(54)

which again requires numerical solution to find \overline{B} . Steepness is given by

$$h = \frac{\frac{\alpha_p(1-r)}{M}(1-e^{-0.2\gamma\overline{B}})}{4 + \frac{\alpha_p(1-r)}{M}(1-e^{-0.2\gamma\overline{B}})}$$
(55)

- which clearly approaches steepness in Eqn 7 as $\gamma \to \infty$.
- Ricker density-dependence and exponential probability of fertilization In this case, \overline{B} satisfies

$$\frac{\alpha_p(1-r)}{M}(1-e^{-\gamma \cdot \overline{B}}) = e^{\beta \cdot \overline{B}}$$
(56)

²⁰¹ and steepness is given by

$$h = 0.2 \left[\frac{1 - e^{-0.2\gamma \cdot \overline{B}}}{1 - e^{-0.2\gamma \cdot \overline{B}}} \right] \left(\frac{\alpha_p (1 - r)}{M} \cdot (1 - e^{-\gamma \cdot \overline{B}}) \right)^{0.8}$$
(57)

²⁰² Numerical Results When Maximum Productivity is Known

To illustrate the above results, we assume that maximum productivity is known and fixed, sex ratio at birth is 0.5, the mean of natural mortality is 0.2 and that in the absence of an Allee effect the steady state biomass is $\overline{B}_0=1000$. For these parameters h = 0.725 for Ricker density dependence and h = 0.556 for Beverton-Holt density dependence. For simplicity, we only present results for cases involving algebraic probability of fertilization We then determine B_c or γ by specifying $p_f(B_0)$ and solving Eqn 25 for B_c ..

Ricker density-dependence and algebraic probability of fertilization 210

The solution of Eqn 52 rapidly converged with Newton's method. In Figure 1a, we show \overline{B} as a 211 function of the as a function of $p_f(B_0)$ as this probability ranges from about 0.6 to 0.999 (note 212 that if $p_f(B_0) = 1$, then B_c must be 0). In Figure 1b we show steepness, given by Eqn 52. 213 We illustrate the Allee effect by plotting the saturating function (left hand side) and exponential 214 function (right hand side) of Eqn 52 and considering their insection point (Figure 1c). As long 215

as the probability of fertilization is less than 1, there will be an Allee value of population size, which is larger for smaller values of probability of fertilization. Populations will decline if their 217 size is smaller than this Allee value. 218

Since steepness is typically computed at 20% of the unfished biomass, we compute the per 219 capita growth rate at 20% of $\overline{B}(p_f(B_0))$ as a function of the probability of fertilization were the 220 population at B_0 (Figure 1d). Note that it is only for relatively low values of the probability of 221 fertilization that per capita growth rate at $0.2\overline{B}$ falls below 0. But, as will be explained below, 222 the Allee effect is important even if per capita growth rate is positive. 223

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Beverton-Holt density dependence and algebraic probability of fertilization 225

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We found the solution of Eqn 50 by using the quadratic formula (and advantage of Beverton-Holt 227 density dependence and algebraic probability of fertilization). In Figures 2a-c we show steepness 228 as a function of the probability of fertilization, the graphical determination of the Allee level 229 (based on rewriting Eqn 50 as $\frac{\alpha_p(1-r)\overline{B}}{\overline{B}+B_c} = (1+\beta \cdot \overline{B})M$), and the per capita growth rate at 20% 230 of \overline{B} , the analogue of Figure 1. Although the numerical values differ, the qualitative results for 231 the two forms of density dependence are very similar, so we focus the rest of this section on 232 Ricker density dependence and algebraic probability of fertilization. 233

The Unstable Steady State is the Wrong Focus of Attention for Population Dynamics with Harvesting

It is common in the discussion of Allee effects to focus on the unstable steady state, because if the population size is below this level, the population will decline even in the absence of humaninduced take. However, when a population is harvested things can go wildly wrong long before the Allee level is reached, as the following analysis suggests.

We assume that in addition to the natural dynamics, which we assume to be Ricker density dependence, the population experiences fishery induced mortality, written either as a fishing mortality rate (F) or as a catch (C)

$$\frac{dB}{dt} = \alpha_p (1-r) \cdot p_f(B) \cdot B \cdot e^{-\beta B} - (F+M) \cdot B$$
(58)

$$\frac{dB}{dt} = \alpha_p (1-r) \cdot p_f(B) \cdot B \cdot e^{-\beta B} - M \cdot B - C$$
(59)

and use two commonly suggested management strategies F = M for Eqn 58 or $C = 0.5M \cdot \overline{B}$ for Eqn 59. We implemented these equations as difference equations, ensuring that B never fell below zero.

The results are shown in Figure 3a for harvest proportional to biomass and Figure 3b for fixed harvest. Perhaps the most important message here is that the Allee effect is hidden but has clear consequences. For example, for probability of fertilization of about 0.74, the per capita growth rate at 20 % of unfished biomass is positive (Figure 1d) but a fishing mortality rate of F = M drives the population to extinction. For a fixed harvest, the result is even more dramatic: even a fertilization probability of almost 90% is insufficient to prevent extinction.

²⁵² The Probability Density for Steepness

As described in Mangel *et al* (2010) in detail, uncertainty in M itself can induce a probability distribution on steepness. That is, all of the equations characterizing steepness are conditioned on a particular value of M. Then if $f_m(m)$ denotes the probability density for M, since the relationship between steepness and rate of mortality is unique, we are easily able to compute the probability density for steepness, $f_h(h)$. All of this can be done for the production model without simulation if α_p is treated as a constant.

For example, we assume that the rate of mortality follows a gamma density with mean 259 0.2 and coefficient of variation $\frac{1}{3}$. This frequency distribution is shown in Figure 4a and the 260 resulting frequency distributions for steepness in Figure 4b for Ricker density dependence and 261 Figure 4c for Beverton-Holt density dependence when the probability of fertilization at \overline{B}_0 is 262 0.887. We note two observations about these figures. First, that for Ricker density dependence, 263 steepness can be arbitrarily large but in this particular case there is little probability of it being 264 larger than about 2.0. Second, although in the absence of Allee effects, steepness for Beverton 265 Holt density dependence cannot fall below 0.2, when there are Allee effects steepness can be less 266 than 0.2. 267

Steepness for the Age-Structured, Two Sex Model with Depen satory SRR

We now turn to the age-structured model. Age structure introduces the biological complexity that the same biomass may be represented by very different age structures, and thus a given biomass does not imply a unique recruitment unless the population is in a deterministic Stable Age Distribution (most likely never; see Wiedenmann *et al* (2009)).

In an age-structured population, we may expect that the probability of successful fertilization depends upon either numbers of biomasses of mature individuals. We will focus on biomass and let B_f and B_m denote the biomass of mature females and males, suppressing the index on time. They are given by

$$B_f = \sum_{a=1}^{a_{max}} N_f(a, t) W_f(a) p_{f,m}(a)$$
(60)

$$B_m = \sum_{a=1}^{a_{max}} N_m(a, t) W_m(a) p_{m,m}(a)$$
(61)

where the terms in the latter equation have the obvious interpretations based on the case without depensation. We assume that the probability of fertilization depends upon the relative mature biomasses. Thus, the analogue of Eqn 25 is

$$p_f(B_f, B_m) = \frac{B_m/B_f}{B_c/B_f + B_m/B_f} = \frac{B_m}{B_c + B_m}$$
(62)

where B_c has exactly the same interpretation as in the production model. The analogue of Eqn 282 26, for the exponential probability of fertilization is

$$p_f(B_f, B_m) = 1 - p_0^{\frac{B_m}{B_f}} = 1 - exp\left[-\frac{B_m}{B_f}|log(p_0)|\right] = 1 - exp\left[-\gamma \frac{B_m}{B_f}\right]$$
(63)

so that γ has exactly the same interpretation as before. Readers who would prefer to think of probability of fertilization in terms of mature numbers rather than mature biomass are encouraged to reproduce the calculations that follow using mature numbers.

A subtlety now arises. Steepness is defined by spawning biomass reduced from its unfished 286 level to 20% of that level. For the production model, this can happen in only one way. However, 287 for the age-structured model there is an infinite number of ways of reducing mature male and 288 female biomasses so that the total is 20% of the unfished level. For example, if we interpret 289 20% of the unfished level to mean that mature male biomass is 20% of its unfished level and 290 mature female biomass is 20% of its unfished level. Then according to Eqn 62 the probability of 291 successful fertilization will decline. However, according to Eqn 63 it will remain the same, which 292 might occur for species in very tight spawning aggregations and highly fecund males. Overall, 293 however, it seems that the algebraic probability of fertilization may capture the effects that we 294 seek to explore more effectively, so we shall use it. 295

²⁹⁶ Beverton-Holt Density Dependence and Algebraic Probability of Fertilization

We begin with Beverton-Holt density dependence and algebraic probability of fertilization for two reasons. First, the calculations are the simplest ones possible – there are no transcendal equations and all quantities can be determined using no more than the quadratic formula. Second, in the absence of Allee effects, steepness for this case ranges between 0.2 and 1, so results are more easily interpreted than for Ricker density dependence.

If density dependence is caused by the entire biomass (not just spawning biomasss) the production of young of the year females and males in the steady state is

$$\overline{N}_f(0) = (1-r)\alpha_s \overline{B}_f \cdot \frac{1}{1+\beta \overline{B}} \cdot \frac{\overline{B}_m}{B_c + \overline{B}_m}$$
(64)

$$\overline{N}_m(0) = r\alpha_s \overline{B}_f \cdot \frac{1}{1+\beta \overline{B}} \cdot \frac{\overline{B}_m}{B_c + \overline{B}_m}$$
(65)

³⁰⁴ and in analogy with Eqn 16 we have

$$\overline{B}_f = \overline{N}_f(0)\overline{W}_f \tag{66}$$

$$\overline{B}_m = \overline{N}_m(0)\overline{W}_m \tag{67}$$

$$\overline{B} = \overline{N}_f(0) < W_f > + \overline{N}_m(0) < W_m >$$
(68)

where $\langle W_f \rangle = \sum_{a=1}^{a_{max}} S_f(a) W_f(a)$ and $\langle W_m \rangle = \sum_{a=1}^{a_{max}} S_m(a) W_m(a)$ are the average masses females and male fish, regardless of the maturation status.

307 Note from Eqns 64 and 65 that

$$\frac{\overline{N}_m(0)}{\overline{N}_f(0)} = \frac{r}{1-r} \tag{69}$$

so that it is helpful to define $\rho_r = \frac{r}{1-r}$ and compactly write $\overline{N}_m(0) = \rho_r \overline{N}_f(0)$.

We now use Eqns 65, 67 and 68 to rewrite Eqn 64 as a single equation for $\overline{N}_f(0)$

$$\overline{N}_f(0) = (1-r)\alpha_s \overline{N}_f(0)\overline{W}_f \cdot \frac{1}{1+\beta \overline{N}_f(0)(\langle W_f \rangle + \rho_r \langle W_m \rangle)} \cdot \frac{\rho_r N_f(0)W_m}{B_c + \rho_r \overline{N}_f(0)\overline{W}_m}$$
(70)

310 whic can be re-arranged to give

$$1 + \beta \overline{N}_f(0) [\langle W_f \rangle + \rho_r \langle W_m \rangle] = \frac{(1 - r)\alpha_s \overline{W}_f \rho_r \overline{N}_f(0) \overline{W}_m}{B_c + \rho_r \overline{N}_f(0) \overline{W}_m}$$
(71)

and cross-multiplying by the denominator on the right hand side, we see that Eqn 71 – as complicated as it looks — is simply a quadratic equation for the single unknown $\overline{N}_f(0)$. Once we find that, we know all the other steady state population numbers and biomasses from Eqns 65-68. Assuming that 20% of unfished biomass is understood as equivalent reductions in female and male populations, steepness is computed from

$$h = \frac{0.2\alpha_s \overline{B}_f \frac{1}{0.2\beta\overline{B}} \frac{0.2B_m}{B_C + 0.2\overline{B}_m}}{\alpha_s \overline{B}_f \frac{1}{\beta\overline{B}} \frac{0.2\overline{B}_m}{B_C + \overline{B}_m}}$$
(72)

317 which simplifies to

$$h = 0.04 \left[\frac{1 + \beta \overline{B}}{1 + 0.2\beta \overline{B}} \left[\frac{B_c + \overline{B}_m}{B_c + 0.2\overline{B}_m} \right]$$
(73)

Note that the left hand side of Eqn 71 is exactly $1 + \beta \overline{B}$, so we solve that equation for $\beta \overline{B}$ and substitute into Eqn 73 to obtain the final result for steepness with Beverton-Holt density dependence and

$$h = 0.2 \left[\frac{(1-r)\alpha_s \overline{W}_f \frac{\rho_r \overline{N}_f(0)\overline{W}_m}{B_c + \rho_r \overline{N}_f(0)\overline{W}_m}}{4 + (1-r)\alpha_s \overline{W}_f \frac{\rho_r \overline{N}_f(0)\overline{W}_m}{B_c + \rho_r \overline{N}_f(0)\overline{W}_m}} \right] \left[\frac{B_c + \rho_r \overline{N}_f(0)\overline{W}_m}{B_c + 0.2\rho_r \overline{N}_f(0)\overline{W}_m} \right]$$
(74)

Note that if we set $B_c = 0$ then we recover Eqn 20 as must happen.

The Deterministic Estimate of Steepness of Bigeye Tuna Thun-*nus obesus*

For this example, we used the same parameters as in Mangel *et al* (2010) and assumed that the rate of mortality, $M_f(a)$ (Eqn 12), for females could be determined from the allometry for fish given by McCoy and Gillooly (2008) assuming a dry mass of 55% of wet mass, and that the rate of mortality for males is given by

$$M_m(a) = \frac{k_m}{k_f} M_f(a) \tag{75}$$

where k_m and k_f are respectively the von Bertalanffy growth rates for males and females respectively. We used exactly deterministic version of the procedure, based on the allometry in McGurk (1986), in Mangel *et al.* (2010) for determining α_s .

The two parameters that remain to be specified are β and B_c . In principle, β can be computed from the dynamics underlying the Beverton-Holt recruitment function, in which per capita larval mortality is a linear function of larval numbers (Mangel 2006, pg 213) but for
illustration here we proceed differently. Note from Eqn 71 that in the absence of Allee effects
we have

$$1 + \beta \overline{N}_f(0)[\langle W_f \rangle + \rho_r \langle W_m \rangle] = (1 - r)\alpha_s \overline{W}_f$$
(76)

The the term in [] on the left-hand side and the right-hand side of Eqn 76 is known once the 336 life history parameters are known. Thus (as observed by many authors), the parametrization 337 of the Beverton-Holt stock-recruitment relationship that we have used means β scales the size 338 of $\overline{N}_{f}(0)$ and that if we specify one of them the other is fixed by the life history parameters. 339 Consequently, one can imagine that both β and $\overline{N}_f(0)$ are measured in some appropriate volume 340 of ocean. For the computations here, we assume that in the absence of Allee effects, $\overline{N}_f(0)$ would 341 be 500 individuals. We then determine β from Eqn 76. When β is determined in this manner, 342 the only unknown in Eqn 53 is the value of $\overline{N}_f(0)$ in the presence of Allee effects; we find this 343 using the quadratic formula. 344

Steepness is then computed from Eqn 74. Clearly steepness depends upon B_c , however as with the production model it is difficult to interpret results in terms of B_c , so we plot steepness as a function of the probability of fertilization, given by

$$p_f = \frac{\rho_r \overline{N}_f(0) \overline{W}_m}{B_c + \rho_r \overline{N}_f(0) \overline{W}_m}$$
(77)

We find that the probability of fertilization is 1 (i.e. $B_c = 0$) steepness is close to 1, but not equal to 1 (Figure 5). However the relationship is highly nonlinear. If $p_f = .976$, then h = 0.908and but if $p_f = 0.9$ of the eggs are fertilized, steepness is about 70%. Note that as with the production model it is possible for steepness to fall below 0.2 if Allee effects are considered.

352 An Improved Approach for Estimating α_s

The stochastic simulation used in Mangel *et al* (2010) is not appropriate for case in which the mechanism of depensation is reduced probability of fertilization, because we must track the size of both male and female populations. In this section, we introduce an improved stochastic
simulation, which can be used for either the depensatory case or the non-depensatory case.

As noted in Mangel *et al* (2010), α_s can be interpreted as the maximum number of new individuals added to the population per unit of spawning biomass before density dependence acts on the recruited class (and α_p can be computed by multiplying α_s by the biomass of a recruit. In the previous paper we used an artifice of 'populations' but here we return to the more common approach based on cohort analysis.

³⁶² The Cohort Based Computation of α_s

We implement the calculations described below in a stochastic simulation over survival trajectories, but for simplicity ignore the index on the iterate of the simulation in the following description. Imagine a cohort of N_0 individuals in which individuals are indexed by *i*. We begin by drawing a random variable U_s which we compare with the sex ratio at birth to determine whether the *i*th fish is a female (s(i) = 1) or a male (s(i)=2). In this manner we determine the number of female and male fish, $N_f(0)$ and $N_m(0)$, in the cohort.

Next we compute the number of females and males of age a using the survival functions, so that $N_f(a) = N_f(0) \cdot S_f(a)$ and $N_m(a) = N_m(0) \cdot S_m(a)$ where $S_f(a)$ and $S_m(a)$ are respectively the probabilities that a female or male survives to age a. In this way we uniquely identify the age a(i) of the i^{th} fish, in which there are a total of $N_T(a) = N_f(a) + N_m(a)$ fish of age a.

Once the age of the i^{th} fish is known we are able to compute the probability that it is mature, $p_m(i)$, from the schedule of maturity. This is a binary variable, with $p_m(i) = 0$ corresponding to an immature fish and $p_m(i) = 1$ corresponding to a mature fish.

In order to account for fertilization based depensation, we must specify the size of the mating group, G. We let B_m and B_f respectively denote the biomass of males and females in the currently simulated mating group and E_T denote the number of eggs that survive to recruit to the population produced by females in the currently simulated mating group. One can imagine a variety of means of assembling the mating group. For example, mature individuals may join mating groups randomly (the pseudocode for this situation is given below) or they may join through size association (so that larger fish are paired with larger fish; smaller fish are paired with smaller fish).

If mating groups involve random association of mature fish, the following pseudocode can befollowed:

Step 1. Set the size of the current group to 0, and set $B_m = B_f = 0$.

Step 2. Draw a test value, i_{test} that is uniformly distributed across the total number of fish. If $p_m(a(i_{test})) = 0$, so that the fish is immature, return to Step 1. If $p_m(a(i_{test})) = 1$, so that the fish is mature, proceed to Step 3.

³⁹⁰ Step 3. Increment the current size of the mating group by 1 individual.

Step 4. If $s(i_{test}) = 2$, so that the test individual is a male, increment B_m by $W_m(a(i_{test}))$.

Step 5. If $s(i_{test}) = 1$, so that the test individual is a female, increment B_f by $W_f(a(i_{test}))$, compute the number of surviving eggs (using the same, but corrected and improved, algorithm as in Mangel *et al* (2010)) and increment total eggs E_T by this amount.

Step 6. If the current group size is less than G return to Step 2. Otherwise, continue to Step 7a or Step 7b.

At this point another decision must be made. If one wished to use Eqn 56, in which a mean probability of fertilization is applied, then

399 Step 7. Compute $\alpha_s = \frac{E_T}{B_f}$.

Step 8. Compute $p_f(B_f, B_m)$ using either Eqn 44 or 45. In this way one obtains the effective maximum production of the test fish.

402 Step 9. If current group size is less than G return to Step 2.

⁴⁰³ Increased Mortality as the Depensatory Mechanism

An alternative depensatory mechanism is that mortality rate increases as population size declines (George Watters, personal communication). For example, we might modify the production model as

$$\frac{dB}{dt} = \alpha_p (1-r)Bg(B) - M_0 \left(\frac{B+B_2}{B}\right)$$
(78)

where g(B) denotes the density dependent component of reproduction, M_0 the rate of natural mortality when biomass is large, and B_2 the value of biomass at which the rate of natural mortality is twice M_0 . For the case of Beverton-Holt density dependence, the steady state biomass satisfies

$$\frac{\alpha_p(1-r)}{M_0} = 1 + \frac{B_2}{\overline{B}} + \beta \overline{B} + \beta B_2 \tag{79}$$

411 which is once again a quadratic equation for \overline{B} .

412 Empirical Assessment for the Depensatory Mechanism

The social psychologist Kurt Lewin is reknowned for his comment that 'there is nothing as practical as a good theory' (Lewin 1951, pg 169). Rothman (2004) revisted Lewin and noted 'Although Lewin may have been right that there is "nothing more practical than a good theory" (p.169; [24]), his dictum rests on the assumption that good theories are available to address practical problems. The development of "good" theories that is, theories that are both accurate and applicable has been hindered by a breakdown in the on-going collaboration between basic and applied behavioral scientists.' (pg 6).

As emphasized in Mangel *et al.* (2010), steepness is a derived quantity – related to things that can be measured, but itself never measured. For the theory of steepness developed here to be practicable, it is necessary to find a way to measure the effect of depensation, captured in either Eqn 25 or 26.

Some very simple statistical methods can be applied to estimate B_c in Eqn 25 or γ in Eqn 26. Inverting Eqn 25 and simplifying shows that (suppressing the dependence of the probability 426 of fertilization on biomass)

$$\frac{1}{p_f} - 1 = B_c \cdot \frac{1}{B} \tag{80}$$

so that B_c can be estimated as the slope of the plot of $\frac{1}{p_f} - 1$ vs. 1/B. (This is similar to methods for estimating the rate constant in Michaelis-Menten enzyme kinetics). Similarly, Eqn 26 can be simply manipulated to give

$$log(1 - p_f) = \gamma \cdot B \tag{81}$$

430 so that γ can be estimated as the slope of the plot of $log(1 - p_f)$ vs B.

Alternatively, Eqns 25 and 26 can be viewed as the foundation of nonlinear statistical models,
 particularly when converted to a logit-form.

Finally, it is natural to consider Bayesian methods by introducing appropriate priors for either B_c or γ . All of this remains to be done.

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Appendix: The Production Model as a System of Stochastic Dif ferential Equations

The gamma density that we use to characterize natural mortality can be viewed as the steady
state frequency distribution for the following stochastic differential equation (Dennis and Costantino
1988, Costantino and Desharnais 1991)

$$dM = M[\lambda - \mu \cdot M]dt + \sigma_M M dW_1 \tag{A1}$$

where λ and μ have their usual interpretations for logistic growth (λ is maximum per capita growth rate and carrying capacity is $\frac{\lambda}{\mu}$), dW_1 is an increment of standard Brownian motion (Mangel 2006) and σ_M is the standard deviation of fluctuations in mortality. Eqn A1 can be interpreted as follows: given that M(t) = m, then dM = M(t+dt) - M(t) is normally distributed with mean

$$E[dM] = m(\lambda - \mu m)dt + o(dt) \tag{A2}$$

487 and variance

$$Var[dM] = \sigma_M^2 m^2 dt + o(dt) \tag{A3}$$

and variance, where o(dt) represents terms that are higher order than dt.

This observation suggests that we can interpret Eqn 1 or the equivalent using Ricker density dependence and the probability density for gamma in the context of stochastic differential equations. If we let g(B) denote the density dependence, then if the source of stochasticity for changes in biomass is a birth and death process, the stochastic version of either equation is(Mangel 1994, 2006)

$$dB = [\alpha_p(1-r)Bg(B)p_f(B) - M \cdot B]dt + \sqrt{\alpha_p(1-r)Bg(B)p_f(B) + M \cdot B} \cdot dW_2$$
(A4)

where dW_2 is another increment in standard Brownian motion. The interpretation for the distribution, mean, and variance of dB = B(t + dt) - B(t) condition on B(t) = b is similar to the one given above for dM.

⁴⁹⁷ If f(m, b) is the stationary probability density for M and B, then under the assumptions in ⁴⁹⁸ Eqns A1-A4, it will satisfy the equation

$$\frac{1}{2} \left[\sigma_M^2 m^2 \frac{\partial^2 f}{\partial m^2} + (\alpha_p (1-r) bg(b) p_f(b) + m \cdot b) \frac{\partial^2 f}{\partial b^2} \right] - \left[m(\lambda - \mu m) \frac{\partial f}{\partial m} \right] - \left[(\alpha_p (1-r) bg(b) p_f(b) - m \cdot b) \frac{\partial f}{\partial b} \right] = 0$$
(A5)

500 Whether or not this proves to be useful remains to be seen.

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516 Captions for Figures

Figure 1. Results for the production model with illustrative parameters. a) The steady state biomass in the presence of an Allee effect as a function of the probability of fertilization $p_f(\overline{B}_0)$ when biomass is $\overline{B}_0 = 1000$; b) Steepness for the production model with Ricker density dependence and algebraic probability of fertilization as a function of the probability of fertilization $p_f(\overline{B}_0)$; c) Graphical determination of the unstable steady state biomass as the solution of the associated transcendental equation. d) Per capita growth rate at 20% of steady state biomass as a function of of the probability of fertilization $p_f(\overline{B}_0)$.

524

Figure 2. Similar calculations for Beverton Holt density dependence and algebraic probability of fertilization. a) Steepness as a function of the probability of fertilization when biomass is 1000, $p_f(\overline{B}_0)$; b) Illustration of graphical determination of the unstable steady state. c) Per capita growth rate at 20% of steady state biomass as a function of of the probability of fertilization when biomass is 1000, $p_f(\overline{B}_0)$.

530

Figure 3 When there is an Allee effect, otherwise sustainable harvesting may become unsustainable. We show, as a function of the probability of fertilization when biomass is 1000, $p_f(\overline{B}_0)$ the biomass trajectory for fishing mortality F = M (panel a) or fixed catch $C = 0.5M\overline{B}$ (panel b)

Figure 4 A probability distribution for natural mortality (panel a; here the gamma density used by Mangel *et al* (2010)) induces a probability distribution for steepness for Ricker density dependence and algebraic probability of fertilization (panel b) or Beverton-Holt density dependence and algebraic probability of fertilization (panel c).

539

Figure 5 The point estimate of steepness for bigeye tuna as a function of the probability that an egg is fertilized when the population is in the stable age distribution.



543 Figure 1a.pdf



544 Figure 1b.pdf



Biomass, B

545 Figure 1c.pdf



546 Figure 1d.pdf



548 Figure 2a.pdf



Biomass, B

549 Figure 2b.pdf



550 Figure 2c.pdf



552 Figure 3a.pdf



Figure 3b.pdf



(a)

Mortality rate, m

555 Figure 4a.pdf



Steepness, h

556 Figure 4b.pdf

Figure 4c.pdf





559 Figure 5.pdf

560 Figure 5. The point estimate of steepness for the age structured model.